Multivariate Bayesian Logistic Regression for Clinical Safety Data

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### Logistic Regression for Subgroup Analyses of Multiple Events

- Start from a Set of *Medically Related* Events to Study
  - Set of ad-hoc events, or all events within a MedDRA SOC
  - Fit Logistic Regressions to each AE as a Response
    - Use exactly the same predictor model for each AE
      - Age, gender, concomitant medication, medical history, etc.
    - Include treatment and interactions with treatment as predictors
    - Generate parameter estimates for predictors and interactions
  - Empirical Bayes Shrinkage of Estimated Coefficients
    - Coefficients of each predictor borrow strength across AEs
    - Overall treatment and interaction effects shrink toward 0



# **Multiple Medically Related Events**

#### Consider Ten MedDRA PTs

- Anuria Dry mouth Hyperkalaemia Micturition urgency Nocturia Pollakiuria Polydipsia Polyuria Thirst Urine output increased
- All seemed somewhat Treatment related in 2x2 analyses

Want Ten Separate Estimates of Treatment Effect

- But some or many of them may have a common cause
  - Common side effects of diuretics
- Analyze them with a common statistical model
- Do they have similar responses to various predictor variables?
- Let the data decide how much they should "borrow strength from each other



### Multivariate Bayesian Logistic Regression (MBLR)

- Use the Same Covariates to Predict all 10 Responses (in Addition to Treatment vs. Placebo Estimates)
  - Sex (F, M)
  - Race (Black, White, Other)
  - Age Group (< 55, 55 to 65, 65 to 75, > 75)
  - Indication (4 Trials w/ Indication 1, 4 Trials with Indication 2)
    - We could have used Trial itself as a predictor instead of Indication
  - Renal Medical History (Yes, No)
- Five Covariates Need 8 Degrees of Freedom



## **Rationale for Use of Covariates**

- Since all Trials Were Randomized, Why Adjust for Covariates? Won't They all Balance Out Anyway?
  - Depending on sample sizes, will not be perfect balance
  - If covariates have strong effects, adjustment for them will reduce residual variance and therefore Treatment effect uncertainty
  - Less focus on a single pre-specified model for safety analyses than for efficacy analyses
  - Main Rationale—Treatment by Covariate Interactions
    - Estimating Treatment x Covariate interactions in a safety analysis is equivalent to searching for vulnerable subgroups
    - MBLR– cross every covariate with the Treatment effect



### **Rationale for EB Model Across Events**

- Coping with Fine Granularity of Adverse Event Data
  - Compare T vs. C on 10 varieties of renal or dehydration issues
  - Approach 1—separate analyses of all 10 events
    - Small counts lead to non significant comparisons
    - Adjustment for multiple comparisons further reduces sensitivity
  - Approach 2—define a single event as union of the 10 events
    - Significant differences may be washed out by the pooling
    - Even if significant, little information about original 10 differences
- Compromise Approach—EB Hierarchical Model
  - 10 individual estimates that "borrow strength" from each other
  - Estimate separate vector of coefficients for each AE
    - But a prior distribution shrinks corresponding coefficients across AEs toward each other
    - The amount of shrinkage is controlled by certain prior variances that are also estimated from the data
    - Treatment-Covariate interaction effects, which are *apriori* less likely, are also shrunk toward the null hypothesis value of 0



### **Defining Regression Effect Estimates**

- Include every Treatment by Covariate Interaction
- Statistical Model for P<sub>ik</sub> = Prob(Event k in i<sup>th</sup> Patient)
  - $X_{ig} = g^{th}$  covariate;  $T_i = Treatment$  arm indicator

$$P_{ik} = 1/[1 + exp(-Z_{ik})]$$

$$Z_{ik} = \alpha_{0k} + \sum_{1 \leq g \leq G} X_{ig} \alpha_{gk} + T_i \left(\beta_{0k} + \sum_{1 \leq g \leq G} X_{ig} \beta_{gk}\right)$$



# **Bayesian Shrinkage Models**

- Statistical Validity of Searching for Extreme Differences
  - Most significant adverse event or patient subgroup
- Classical Approach to Post-Hoc Interval Estimates
  - Maintain centers of CI at observed differences
  - Expand widths of every CI
  - Expansion is greater the more differences you look at
    - If you look at too many, the CI's are too wide to be useful
  - Bayesian Approach
    - Requires a prior distribution for differences
      - Can estimate it from the multiple observed differences available
    - Centers of CI's are "shrunk" toward average or null difference
      - High-variance differences shrink the most
    - Widths of Cl's usually shrink a little too
      - The more you look at, the better you can model the prior dist.



## **Prior Distributions for Coefficients**

#### Two-Stage Hierarchical Model

- Covariate main effects  $\alpha_{qk}$  shrink toward means across issues
- Treatment main effects β<sub>0k</sub> shrink toward each other
- Treatment interactions  $\beta_{qk}$  also shrink toward 0
- Four prior standard deviations control amount of shrinkage

- Let  $\phi = (\sigma_A, \sigma_0, \sigma_B, \tau)$ ; prior distributions uniform (0, d = 1.5)

- { $\alpha_{0k}$ , A<sub>g</sub>, B<sub>0</sub>} have uniform priors (- $\infty$ , + $\infty$ )
- Remaining parameters have prior distributions:
  - $$\begin{split} \alpha_{gk} | A_g &\sim N(A_g, \, \sigma_A^{\ 2}) \,, & k = 1, \, \dots, \, K \ ; \ g = 1, \, \dots, \, G \\ \beta_{0k} | B_0 &\sim N(B_0, \, \sigma_0^{\ 2}) \,, & k = 1, \, \dots, \, K \\ \beta_{gk} | B_g &\sim N(B_g, \, \sigma_B^{\ 2}) \,, & k = 1, \, \dots, \, K \ ; \ g = 1, \, \dots, \, G \\ B_g &\sim N(0, \, \tau^2) \,, & g = 1, \, \dots, G \end{split}$$



# **Computational Approach**

- Prefer Not To Use MCMC Methodology
  - Commercial software designed for non-statistician users
  - Convergence and non-exact-repeatability are issues
  - Scale-up problem: several hundred regression parameters
- Approximate Posterior Distributions
  - Discrete approximation of posterior of  $\phi = (\sigma_A, \sigma_0, \sigma_B, \tau)$

– { $\pi_s$ ,  $\phi_s$  s = 1, ..., S} defines S-point discrete distribution,  $\Sigma_s \pi_s = 1$ 

- Normal approximation to  $P(\theta | data, \phi)$ 
  - $\theta = (A_1, ..., A_G, B_0, ..., B_G, \alpha_{01}, ..., \alpha_{GK}, \beta_{01}, ..., \beta_{GK})$
  - Can use modified logistic regression likelihood for P( $\theta \mid data, \phi$ )
  - Log P( $\theta$  |data,  $\phi_s$ ) concave and easy to maximize, ~ N( $\theta_s$ , V<sub>s</sub>)
- Posterior of θ Approximately N(μ, V)
  - $\mu = \Sigma_s \pi_s \theta_s$   $V = \Sigma_s \pi_s [V_s + (\theta_s \mu)(\theta_s \mu)^t]$

# **Computing** $\{\pi_s, \phi_s \mid s = 1, ..., S\}$

- Density P( $\phi_s$  |data)  $\approx$  g( $\phi_s$ )  $\propto$  P(data |  $\phi_s$ ,  $\theta_s$ ) det(V<sub>s</sub>)<sup>1/2</sup>
  - Maximized likelihood × (approx. factor for integrating out  $\theta$ )
- Steepest Ascent (Numerical Derivs) to Maximize g(\u00f3)
- Construct Response Surface Design Around Maximum
  - 16 point central composite design at each of two radii
  - Fit 4-D quadratic response surface model (rsm) to log  $g(\phi)$
  - Use fitted surface to rescale the 33-point design and refit rsm
- Adjust Final 33 Values of  $\pi_s$  so that Means and Variances of Discrete Dist. Match Continuous Estimates from RSM
  - Minimize K-L =  $\Sigma_s g(\phi_s) \log(g(\phi_s)/\pi_s)$  subject to constraints
- **Complete Estimation**  $\approx$ 10 Seconds if  $\theta$  Has 200 Elements



# **Comparing MBLR to "Standard" LR**

- Logistic Regression on Rare Events with Several Covariates and Interactions Can Often Fail to Get Reasonable Answers
  - Certain combinations of covariates seem to predict perfectly, leading to coefficient estimates that diverge to + or – infinity
  - Related terms: Separation, Sparsity, Nonidentifiability
  - Gelman et al (2008 Annals of Applied Statistics)
    - Suggests using a very weak prior distribution on the coefficients to get more reasonable answers and prevent divergence
    - Calls method Regularized Logistic Regression (RLR)
- Comparisons of MBLR to RLR
  - RLR: same model as MBLR except that  $\sigma_A = \sigma_0 = \sigma_B = \tau = 5$
  - Typically, MBLR estimates of prior standard deviations < 1</li>



# **Back to the Example**

Treatment	Comparator	Chi	Shrunken
Events	Events	Statistic	Odds Ratio
8	0	2.61	9.02
308	65	11.42	4.31
218	162	1.34	1.15
13	3	2.18	3.10
19	7	1.95	2.17
193	34	9.55	4.98
49	4	5.63	8.88
100	17	6.89	4.93
543	66	18.38	8.16
13	1	2.92	6.45
Treatmer	nt = 3,110	Comparate	or = 2,642
	Treatment   Events   8   308   218   13   19   193   49   100   543   13   Treatment	Treatment   Comparator     Events   Events     8   0     308   65     218   162     13   3     19   7     193   34     49   4     100   17     543   66     13   1     Treatment = 3,110   1	Treatment   Comparator   Chi     Events   Events   Statistic     8   0   2.61     308   65   11.42     218   162   1.34     13   3   2.18     19   7   1.95     193   34   9.55     49   4   5.63     100   17   6.89     543   66   18.38     13   1   2.92     Treatment = 3,110   Comparator

Statistics for 10 Issues Related to Dehydration/Renal Function for a Pool of 8 Trials



# **Covariate Patient Counts**

	Treatment	Comparator	
Gender=F	908	685	
Gender=M	2202	1957	
Race=Black	323	216	
Race=White	2516	2228	
Race=Other	271	198	
Indication=1	796	347	
Indication=2	2314	2295	
Renal History=Y	190	191	
Renal History=N	2920	2451	
Age=50 or Under	382	348	
Age=51 to 65	1089	902	
Age=66 to 75	948	820	
Age=Over 75	691	572	
All Patients	3110	2642	

**Distribution of Patients by Covariates and Treatment Arm** 



### **Treatment Effects: RLR vs MBLR**

Odds Ratio Estimates for TERM = Treatment



Odds Ratio (90% CI)



RESPONSE. METHOD

## **Robustness to Post-Hoc Selection**

- Simulation Study of Bayesian Estimation
  - Draw "true parameters" from the prior distributions 1000 times
  - Estimate main and interaction effects each time
    - Get both MBLR and RLR estimates

Focus on Estimating the "Most Significant" Interaction

- 80 Interactions (8 covariates x 10 response events)
- For each simulation, select  $\beta_{gk}$  that has *largest*  $b_{gk}/se_{gk}$
- Compare accuracy of estimates and confidence limits

	SIM.COEF	SD.SIMC	BIAS	RMSE	Z.SCORE	CI.05	CI.95
MBLR	1.7651	0.6094	0.0005	0.2923	-0.0052	0.067	0.056
RLR	1.7445	0.5981	0.2184	0.4330	0.5794	0.008	0.135



# **Safety Analyses of Clinical Data**

Analysis of Drug Trial Adverse Event Data Is Challenging

- Small event counts since trials are sized for efficacy
- Multiple comparisons issues
- Combined Analyses of Multiple Trials Is Important
  - CDISC data standards make pooling data easier
  - This is a form of pooled-data meta-analysis
- Multivariate Bayesian Logistic Regression (MBLR)
  - Multivariate estimation of many possibly medically related AEs
  - Borrowing strength as a solution to the granularity problem
  - Search for vulnerable subgroups involves post-hoc selection
  - Bayesian shrinkage provides multiple-comparisons robustness

